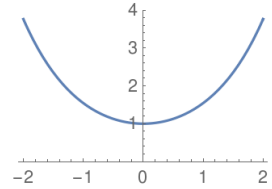


## SAMPLE SOLUTIONS EXERCISE 6

### EXERCISE 6.1: TRANSFORMING PROBABILITY DENSITIES

(6P)

Consider the curve  $y = \cosh x$ . The aim of this exercise is to decorate this curve in the interval  $x \in [x_1, x_2]$  with random points in such a way that the density of the points is uniform along the curve. Since the slope of the curve varies, this means that the  $x$ -coordinates of these points are *not* uniformly distributed.



- (a) Compute the probability density  $p(x)$  of the  $x$ -coordinates. Hint: the density of the points per arc length of the curve has to be constant. (2P)
- (b) Normalize the probability density on the interval  $x \in [x_1, x_2]$ . (1P)
- (c) Find a function  $f : z \mapsto x = f(z)$  such that it maps a uniform probability density  $p(z) = \text{const}$  to the non-uniform probability density  $p(x)$  calculated in (a)-(b). (2P)
- (d) Adjust the integration constant in (c) in such a way that  $f$  maps  $[z_1, z_2] \mapsto [x_1, x_2]$  with  $f(z_1) = x_1$  and  $f(z_2) = x_2$ . Specialize the result for a standard random number ( $z_1 = 0, z_2 = 1$ ). (1P)

### SAMPLE SOLUTION

- (a) The infinitesimal line element  $d\ell$  of the curve  $y = f(x) = \cosh x$  is given by (1P)

$$d\ell = \sqrt{1 + (f'(x))^2} dx = \sqrt{1 + \sinh^2(x)} dx = \cosh(x) dx$$

We start with the ansatz

$$|p(\ell) d\ell| = |p(x) dx|$$

where  $p(\ell) = C$  is constant along the curve. Using this ansatz we get (1P)

$$p(x) = C \left| \frac{d\ell}{dx} \right| = C \cosh x.$$

- (b) The normalization can be computed straight-forwardly:

$$\int_{x_1}^{x_2} p(x) dx = C(\sinh x_2 - \sinh x_1) \Rightarrow C = \frac{1}{\sinh x_2 - \sinh x_1}$$

giving

$$p(x) = \frac{\cosh x}{\sinh x_2 - \sinh x_1}.$$

- (c) Again we start with

$$|p(z) dz| = |p(x) dx|,$$

where  $z \in [z_1, z_2]$  is uniformly distributed, meaning that  $p(z) = \frac{1}{z_2 - z_1}$  is constant. Again we get a first-order differential equation

$$\frac{dz}{dx} = z'(x) = \frac{p(x)}{p(z)} = \frac{z_2 - z_1}{\sinh x_2 - \sinh x_1} \cosh x,$$

hence

$$z(x) = \frac{z_2 - z_1}{\sinh x_2 - \sinh x_1} \sinh x + \tilde{C},$$

where  $\tilde{C}$  is the integration constant. However, what we need is not  $z(x)$  but the inverse function  $x(z)$ :

$$x(z) = \operatorname{arcsinh} \left[ \frac{\sinh x_2 - \sinh x_1}{z_2 - z_1} (z - \tilde{C}) \right]$$

(d) Assume that  $f$  maps  $[z_1, z_2]$  onto  $[x_1, x_2]$  with

$$x(z_1) = x_1, \quad x(z_2) = x_2.$$

It is easy to see that the difference  $\sinh(x(z_2)) - \sinh(x(z_1)) = \sinh(x_2) - \sinh(x_1)$  is always true, hence both equations are not independent. Taking one of them we can compute the integration constant

$$\tilde{C} = \frac{z_2 \sinh x_1 - z_1 \sinh x_2}{\sinh x_2 - \sinh x_1}.$$

Plugging that in we arrive at the final result

$$x(z) = \operatorname{arcsinh} \left[ \frac{(z - z_1) \sinh x_2 - (z - z_2) \sinh x_1}{z_2 - z_1} \right].$$

For a standard random number generator  $z \in [0, 1]$  with  $z_1 = 0$  and  $z_2 = 1$  this specializes to

$$x(z) = \operatorname{arcsinh} \left[ \sinh x_1 + z(\sinh x_2 - \sinh x_1) \right].$$

## EXERCISE 6.2: RELATIVE ENTROPY

(2P)

The relative entropy  $H(p||q)$  of two probability distributions  $p_1, \dots, p_N$  and  $q_1, \dots, q_N$ , which is also known as Kullback-Leibler-Divergence  $D(p||q)$ , measures how different the probability distributions are. It is defined by

$$H(p||q) = D(p||q) = \sum_i p_i \ln \frac{p_i}{q_i}.$$

- (a) Use Jensen's inequality to show that  $H(p||q) \geq 0$ . (1P)
- (b) Show that  $H(p||p) = 0$  and  $H(p||u) = \ln N - H(p)$ , where  $u$  is the uniform (constant) probability distribution. (1P)

## SAMPLE SOLUTION

- (a) Since  $-\ln x$  is a convex function, we can apply Jensen's inequality as discussed in the lecture:

$$\begin{aligned} H(p||q) &= \sum_i p_i \ln \frac{p_i}{q_i} = - \sum_i p_i \ln \frac{q_i}{p_i} = \left\langle - \ln \frac{q_i}{p_i} \right\rangle \\ &\geq - \ln \left\langle \frac{q_i}{p_i} \right\rangle = - \ln \left( \sum_i p_i \frac{q_i}{p_i} \right) = - \ln \left( \underbrace{\sum_i q_i}_{=1} \right) = 0. \end{aligned}$$

- (b)

$$\begin{aligned} H(p||p) &= \sum_i p_i \ln \frac{p_i}{p_i} = 0. \\ u_i = 1/N \quad \Rightarrow \quad H(p||u) &= \sum_i p_i \ln \frac{p_i}{1/N} = \underbrace{\sum_i p_i \ln N}_{=1} + \underbrace{\sum_i p_i \ln p_i}_{=-H(p)} = \ln N - H(p). \end{aligned}$$

### EXERCISE 6.3: ENTROPY OF OVERLAPPING PROBABILITY DENSITIES (4P)

Let  $p(x)$  be a given continuously differentiable probability density. Let us shift it to the right and to the left by  $x \rightarrow x \pm a$  and create a mixed probability density of the form

$$q(x) := \frac{1}{2}(p(x-a) + p(x+a)).$$

The purpose of this exercise is that the entropy is minimal for  $a = 0$  where both of them match, that is, the entropy is minimized in the case of a perfect overlap.

- (a) Show that  $\int_{-\infty}^{+\infty} (p'(x+a) - p'(x-a)) dx = 0 \quad \forall a$ . (1P)
- (b) Prove that the entropy  $H_a = - \int_{-\infty}^{+\infty} q(x) \ln q(x) dx$  is extremal for  $a = 0$ . (1P)
- (c) Show that the extremum at  $a = 0$  is a minimum. You can assume that  $p(x)$  and all its derivatives vanish at  $x = \pm\infty$ . (2P)

**Note:** Is this minimization of overlapping functions useful? Yes, it is, you can use it to tune your piano, see <http://piano-tuner.org>. Available for Windows, Mac, Android, and iOS.

## SAMPLE SOLUTION

- (a)

$$\begin{aligned} \int_{-\infty}^{+\infty} (p'(x+a) - p'(x-a)) dx &= \int_{-\infty}^{+\infty} \frac{d}{da} (p(x+a) + p(x-a)) \\ &= \frac{d}{da} \left[ \int_{-\infty}^{+\infty} p(x+a) dx + \int_{-\infty}^{+\infty} p(x-a) dx \right] = \frac{d}{da} (1 + 1) = 0 \end{aligned}$$

(b) The entropy reads:

$$H_a = - \int_{-\infty}^{+\infty} \frac{1}{2} \left( p(x-a) + p(x+a) \right) \ln \frac{1}{2} \left( p(x-a) + p(x+a) \right).$$

As a necessary condition for an extremal point,  $\frac{d}{da} H_a = 0$  has to vanish. Carrying out the derivative we obtain

$$\frac{d}{da} H_a = \frac{1}{2} \int_{-\infty}^{+\infty} \left[ 1 + \ln \left( \frac{1}{2} \left( p(x-a) + p(x+a) \right) \right) \right] \cdot \left( p'(x+a) - p'(x-a) \right) dx = 0$$

Clearly, the round bracket vanishes for  $a = 0$ , hence we have an extremum here.

(c) A straight-forward calculation yields

$$\frac{d^2}{da^2} H_a \Big|_{a=0} = - \int_{-\infty}^{+\infty} (1 + \ln p(x)) p''(x) dx$$

We integrate this expression by parts:

$$\frac{d^2}{da^2} H_a \Big|_{a=0} = - \underbrace{\left[ (1 + \ln p(x)) p'(x) \right]_{-\infty}^{+\infty}}_{=0 \text{ since } p'(\pm\infty)=0} + \int_{-\infty}^{+\infty} \underbrace{\frac{1}{p(x)}}_{\geq 0} \underbrace{p'(x)p'(x)}_{\geq 0} dx > 0$$

Hence we have a minimum.

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( $\Sigma = 12P$ )